Basic Linear Regression

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Basic Linear Regression

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Introduction

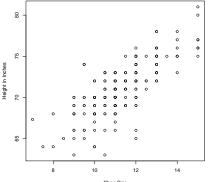
- In this module, we discuss an extremely important technique in statistics Linear Regression.
- Linear regression is very closely related to correlation, and is extremely useful in a wide range of areas.

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Introduction

We begin by recalling our data relating height to shoe size and drawing the scatterplot for the male data.

```
> all.heights <- read.csv("shoesize.csv")
> male.data <- all.heights[all.heights@Gender == "N", ] #Select males
> attach(male.data) #Make Variables Available
> # Draw scatterplot
> plot(Size, Height, lab = "Shoe Size", ylab = "Height in Inches")
```



Shoe Size

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Introduction

The correlation is an impressive 0.77. But how can we characterize the relationship between shoe size and height?

> cor(Size, Height)

[1] 0.7677

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Introduction

- If data are scattered around a straight line, then the relationship between the two variables can be thought of as being represented by that straight line, with some "noise" or error thrown in.
- We know that the correlation coefficient is a measure of how well the points will fit a straight line. But *which straight line is best*?

Introduction

- The key to understanding this is to realize the following:
 - Any straight line can be characterized by just two parameters, a *slope* and an *intercept*, and the equation for the straight line is Y = bX + a, where b is the slope and a is the intercept.
 - Any point can be characterized *relative to a particular line* in terms of two quantities: (a) where its X falls on a line, and (b) how far its Y is from the line in the vertical direction.
- Let's examine each of these preceding points.

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Characteristics of a Straight Line

- Your textbook uses the notation Y = bX + a for a straight line. But there are many different notations, and it will be up to you to keep track of what symbols are used for the slope and intercept!
- For example, for reasons that become apparent very quickly if you take a graduate course, many authors prefer a subscripted notion of the form $Y = \beta_1 X + \beta_0$ in the context of linear regression. In that notation, β_1 is the slope and β_0 is the intercept.

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Characteristics of a Straight Line

- The key point is that the slope is multiplied by X, and so any change in X is multiplied by the slope and passed on to Y. Consequently, the slope represents "the rise over the run," the amount by which Y increases for each unit increase in X.
- The intercept is, of course, the value of Y when X = 0.
- So if you have the slope and intercept, you have the line.

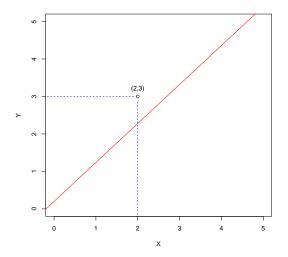
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Characteristics of a Straight Line

- Suppose we draw a line any line in a plane.
- Then consider a point any point with respect to that line.
- What can we say? Let's use a concrete example.
- Suppose I draw the straight line whose equation is Y = 1.04X + 0.2 in a plane, and then plot the point (2,3) by going over to 2 on the X-axis, then up to 3 on the Y-axis.

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Characteristics of a Straight Line



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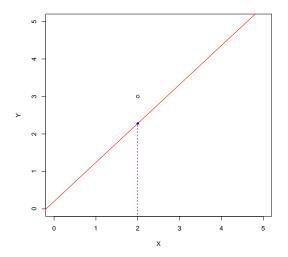
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Characteristics of a Straight Line

- Now suppose I were to try to use the straight line to predict the Y value of the point only from a knowledge of the X value of that point.
- The X value of the point is 2. If I substitute 2 for X in the formula Y = 1.04X + 0.2, I get Y = 2.28.
- This value lies on the line, directly above X. I'll draw that point on the scatterplot in blue.

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Characteristics of a Straight Line



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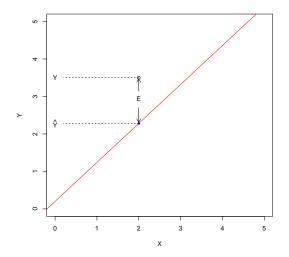
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Characteristics of a Straight Line

- The Y value for the blue point is called the "predicted value of Y," and is denoted \hat{Y} .
- Unless the actual point falls on the line, there will be some error in this prediction. The error is the discrepancy in the vertical direction from the line to the point.

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Characteristics of a Straight Line



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Regression Notation

- Now, let't generalize!
- We have just shown that, for any point with coordinates (X_i, Y_i) , relative to any line Y = bX + a, I may write

$$\hat{Y}_i = bX_i + a$$
 (1)

and

$$Y_i = \hat{Y} + E_i \tag{2}$$

- But we are not looking for *any* line. We are looking for the *best* line. And we have many points, not just one.
- And, by the way, what is the best line, and how do we find it?

The Least Squares Solution

- It turns out, there are many possible ways of characterizing how well a line fits a set of points.
- However, one approach seems quite reasonable, and has many absolutely beautiful mathematical properties.
- This is the *least squares criterion* and the *least squares solution* for a and b.

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The Least Squares Solution

- The least squares criterion states, the best-fitting line for a set of points is that line which minimizes the sum of squares of the E_i for the entire set of points.
- Remember, the data points are there, plotted in the plane, nailed down, as it were. The only thing free to vary is the line, and it is characterized by just two parameters, the slope and intercept.
- For any slope *b* and intercept *a* I might choose, I can compute the sum of squared errors. And for any data set, the sum of squared errors is uniquely defined by that slope and intercept.
- The sum of squared errors is thus a *function* of *a* and *b*.
- What we really have is a problem in minimizing a function of two unknowns.
- This is a routine problem in first-year calculus. We won't go through the proof of the least squares solution, we'll simply give you the result.

The Least Squares Solution

• The solution to the least squares criterion is as follows

$$b = r_{y,x} \frac{s_y}{s_x} = \frac{s_{y,x}}{s_x^2}$$
(3)

and

$$a = M_y - bM_x \tag{4}$$

- Note: If X and Y are both in Z score form, then $b = r_{y,x}$ and a = 0.
- Thus, once we remove the metric from the numbers, the very intimate connection between correlation and regression is revealed!

Creating a Fit Object

- We could easily construct the slope and intercept of our regression line from summary statistics.
- But R actually has a facility to perform the entire analysis very quickly and automatically.
- You begin by producing a *linear model fit object* with the following syntax.

```
> fit.object <- lm(Height ~ Size)</pre>
```

- R is an *object oriented language*. That is, objects can contain data and when general functions are applied to an object, the object "knows what to do."
- We'll demonstrate on the next slide.

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Examining Summary Statistics

• R has a generic function called summary. Look what happens when we apply it to our fit object.

```
> summary(fit.object)
Call:
lm(formula = Height ~ Size)
Residuals:
  Min
      10 Median 30 Max
-7.289 -1.112 0.066 1.356 5.824
Coefficients:
           Estimate Std. Error t value Pr(>|t|)
(Intercept) 52.5460 1.0556 49.8 <2e-16 ***
                      0.0928 17.7 <2e-16 ***
Size
        1.6453
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 2.02 on 219 degrees of freedom
Multiple R-squared: 0.589, Adjusted R-squared: 0.588
F-statistic: 314 on 1 and 219 DF, p-value: <2e-16
```

Examining Summary Statistics

• The coefficients for the intercept and slope are perhaps the most important part of the output. Here we see that the slope of the line is 1.6453 and the intercept is 52.5460.

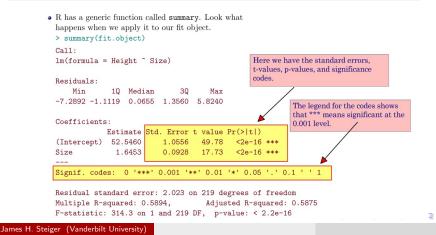
```
    R has a generic function called summary. Look what

 happens when we apply it to our fit object.
 > summary(fit.object)
 Call:
 lm(formula = Height ~ Size)
 Residuals:
      Min
               10 Median
                                30
                                       Max
 -7.2892 -1.1119 0.0655 1.3560 5.8240
                                               These are the estimates for the intercept
                                               and slope of the straight line
 Coefficients:
              Estimate Std. Error t value Pr(>|t|)
 (Intercept) 52.5460
                           1.0556
                                     49.78
                                              <2e-16 ***
 Size
                1.6453
                           0.0928
                                     17.73
                                             <2e-16 ***
 Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
 Residual standard error: 2.023 on 219 degrees of freedom
 Multiple R-squared: 0.5894,
                                     Adjusted R-squared: 0.5875
 F-statistic: 314.3 on 1 and 219 DF, p-value: < 2.2e-16
```

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Examining Summary Statistics

• Along with the estimates themselves, the program provides estimated standard errors of the coefficients, along with *t* statistics for testing the hypothesis that the coefficient is zero.



Examining Summary Statistics

- The program prints the R^2 value, also known as the coefficient of determination. When there is only one predictor, as in this case, the R^2 value is just $r_{x,y}^2$, the square of the correlation between height and shoe size.
- The "adjusted R^{2} " value is an approximately unbiased estimator. With only one predictor, this can essentially be ignored, but with many predictors, it can be much lower than the standard R^{2} estimate.

• The *F*-statistic tests that
$$R^2 = 0$$

• When there is only one predictor, it is the square of the *t*-statistic for testing that $r_{x,y} = 0$.

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Examining Summary Statistics

```
Call:
lm(formula = Height ~ Size)
Residuals:
   Min
       10 Median 30
                                 Max
-7.2892 -1.1119 0.0655 1.3560 5.8240
Coefficients:
                                                            The F
                                                            the A
           Estimate Std. Error t value Pr(>|t|)
                                                            R-squ
(Intercept) 52.5460 1.0556 49.78 <2e-16 ***
                                                            F-tes
Size 1.6453 0.0928 17.73 <2e-16 ***
                                                            hypot
                                                            R-squ
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 2.023 on 219 degrees of freedom
Multiple R-squared: 0.5894, Adjusted R-squared: 0.5875
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```

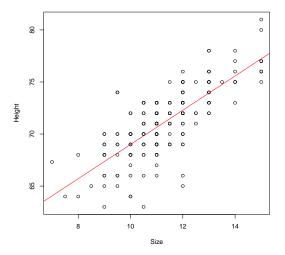
Drawing the Regression Line

- Now we draw the scatterplot with the best-fitting straight line.
- Notice how we draw the scatterplot first with the plot command, then draw the regression line in red with the abline command.

```
> # draw scatterplot
> plot(Size, Height)
> # draw regression line in red
> abline(fit.object, col = "red")
```

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Drawing the Regression Line



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Computing a Predicted Value

- We can now use the regression line to estimate a male student's height from his shoe size.
- Suppose a student's shoe size is 13. What is his predicted height?

 $\hat{Y} = bX + a = (1.6453)(13) + 52.5460 = 73.9349$

- The predicted height is a bit less than 6 feet 2 inches.
- Of course, we know that not every student who has a size 13 show will have a height of 73.93. Some will be taller than that, some will be shorter. Is there something more we can say?

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Thinking about Residuals

- The predicted value $\hat{Y} = 73.93$ actually represents the average height of people with a shoe size of 13.
- According to the most commonly used linear regression model, people with a shoe size of 13 actually have a normal distribution with a mean of 73.93, and a standard deviation called the "standard error of estimate."
- This quantity goes by several names, and in R output is called the "residual standard error."
- An estimate of this quantity is included in the R regression output produced by the summary function.

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Thinking about Residuals

```
Call:
 lm(formula = Height ~ Size)
 Residuals:
     Min 1Q Median 3Q
                                    Max
 -7.2892 -1.1119 0.0655 1.3560 5.8240
 Coefficients:
             Estimate Std. Error t value Pr(>|t|)
 (Intercept) 52.5460 1.0556 49.78 <2e-16 ***
 Size
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                0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
 Signif. codes:
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Thinking about Residuals

• In the population, the standard error of estimate is calculated from the following formula

$$\sigma_e = \sqrt{1 - \rho_{x,y}^2} \, \sigma_y \tag{5}$$

 In the sample, we estimate the standard error of estimate with the following formula

$$s_e = \sqrt{\frac{n-1}{n-2}} \sqrt{1 - r_{x,y}^2} s_y$$
 (6)

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An Example

- Residuals can be thought of as "The part of Y that is left over after that which can be predicted from X is partialled out."
- This notion has led to the concept of *partial correlation*.
- Let's introduce this notion in connection with an example.
- Suppose we gathered data on house fires in the Nashville area over the past month. We have data on two variables — damage done by the fire, in thousands of dollars (Damage) and the number of fire trucks sent to the fire by the fire department (Trucks).
- Here are the data for the last 10 fires.

An Example

	Trucks	Damage
1	0	8
2	0	9
3	1	33
4	1	38
5	1	27
6	2	70
7	2	94
8	2	83
9	3	133
10	3	135

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An Example

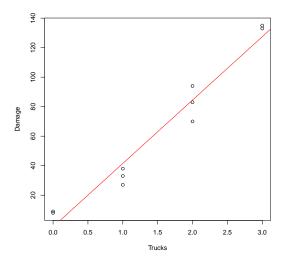
• Plotting the regression line, we see that there is indeed, a strong linear relationship between the number of fire trucks sent to a fire, and the damage done by the fire.

> plot(Trucks, Damage)

> abline(lm(Damage ~ Trucks), col = "red")

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An Example



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An Example

- The correlation between Trucks and Damage is 0.9779.
- Does this mean that the damage done by fire can be reduced by sending fewer trucks?
- Of course not. It turns out that the house fire records include another piece of information. Based on a complex rating system, each housefire has a rating based on the size of the conflagration. These ratings are in a variable called FireSize.
- On purely substantive and logical grounds, we might suspect that rather than fire trucks causing the damage, that this third variable, FireSize, causes both more damage to be done and more fire trucks to be sent.
- How can we investigate this notion statistically?

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An Example

- Suppose we predict Trucks from FireSize. The residuals represent that part of Trucks that isn't attributable to Firesize. Call these residuals *E*_{Trucks}•FireSize</sub>.
- Then suppose we predict Damage from Firesize. The residuals represent that part of Damage that cannot be predicted from FireSize. Call these residuals *E*_{Damage•Firesize}.
- The correlation between these two residual variables is called *the partial correlation* between Trucks and Damage with FireSize partialled out, and is denoted *r*_{Trucks,Damage•FireSize}.

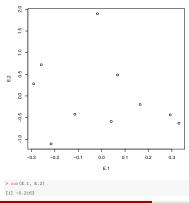
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An Example

- . There are several ways we can compute this partial correlation.
- One way is to compute the two residual variables discussed above, and then compute the correlation between them.



> plot(E.1, E.2)



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An Example

• Another way is to use the textbook formula

$$r_{x,y\bullet z} = \frac{r_{x,y} - r_{x,z}r_{y,z}}{\sqrt{(1 - r_{x,z}^2)(1 - r_{x,z}^2)}}$$
(7)

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[1] -0.2163

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An Example

Partial Correlation

An Example

- The partial correlation is -0.216.
- Once size of fire is accounted for, there is a negative correlation between number of fire trucks sent to the fire and damage done by the fire.

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